

Choquistic Regression: Generalizing Logistic Regression Using the Choquet Integral

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- **Contribution:**

We introduce a new method for (probabilistic) **binary classification**, called **choquistic regression**, which generalizes conventional logistic regression and takes advantage of the **Choquet integral** as a flexible and expressive aggregation operator.

- **Outline:**

- (1) Background on **logistic regression**
- (2) Generalization to **choquistic regression**
- (3) First **experimental results**

- **Logistic regression** modifies linear regression for the purpose of predicting (probabilities of) a **binary class label** instead of real-valued responses.
- The basic model:

$$\begin{aligned} \text{log-odds ratio} \rightarrow \log \left(\frac{\mathbf{P}(y = 1 | \mathbf{x})}{\mathbf{P}(y = 0 | \mathbf{x})} \right) &= w_0 + \sum_{i=1}^m w_i \cdot x_i \\ &= w_0 + \mathbf{w}^\top \mathbf{x}, \end{aligned}$$

where

 linear function of predictor variables

- $\mathbf{x} = (x_1, x_2, \dots, x_m)^\top \in \mathbb{R}^m$ is an instance to be classified,
- $\mathbf{w} = (w_1, w_2, \dots, w_m)^\top \in \mathbb{R}^m$ is a vector of regression coefficients,
- $w_0 \in \mathbb{R}$ is a constant bias (the intercept).

- Equivalently, this can be expressed in terms of **posterior probabilities**:

$$\mathbf{P}(y = 1 | \mathbf{x}) = \left(1 + \exp(-w_0 - \mathbf{w}^\top \mathbf{x})\right)^{-1}$$

$$\mathbf{P}(y = 0 | \mathbf{x}) = 1 - \mathbf{P}(y = 1 | \mathbf{x})$$

- Predictions** are typically made using the following decision rule:

$$\hat{y} = \begin{cases} 0 & \text{if } \mathbf{P}(y = 1 | \mathbf{x}) < 1/2 \\ 1 & \text{if } \mathbf{P}(y = 1 | \mathbf{x}) \geq 1/2 \end{cases}$$

- The parameters of the model (bias, regression coefficients) can be obtained through **Maximum Likelihood (ML)** estimation.
- Given a sample of i.i.d. data

$$\mathcal{D} = \left\{ (\mathbf{x}^{(i)}, y^{(i)}) \right\}_{i=1}^n \subset (\mathbb{R}^m \times \{0, 1\})^n ,$$

the likelihood function is given by

$$\prod_{i=1}^n \mathbf{P} \left(y = y^{(i)} \mid \mathbf{x}^{(i)} \right) ,$$

and the **ML estimate** is the maximizer of (the log of) this function:

$$(\hat{w}_0, \hat{\mathbf{w}}) = \arg \max_{(w_0, \mathbf{w})} \sum_{i=1}^n y^{(i)} \log \theta^{(i)}(w_0, \mathbf{w}) + (1 - y^{(i)}) \log (1 - \theta^{(i)}(w_0, \mathbf{w}))$$

with

$$\theta^{(i)}(w_0, \mathbf{w}) = \left(1 + \exp(-w_0 - \mathbf{w}^\top \mathbf{x}^{(i)}) \right)^{-1}$$

- Logistic regression is very popular and widely used in practice.
- It is **comprehensible** and easy to **interpret**, especially since the influence of each variable can easily be captured from the model:

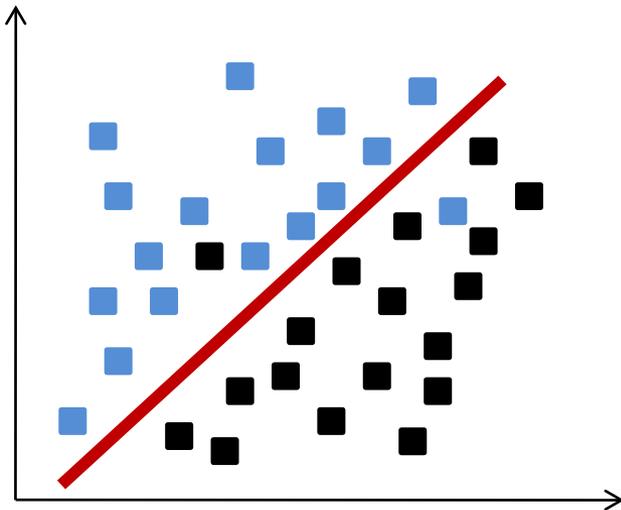
$$\log \left(\frac{\mathbf{P}(y = 1 | \mathbf{x})}{\mathbf{P}(y = 0 | \mathbf{x})} \right) = w_0 + \boxed{w_1} \cdot x_1 + w_2 \cdot x_2 + \dots + w_m \cdot x_m$$



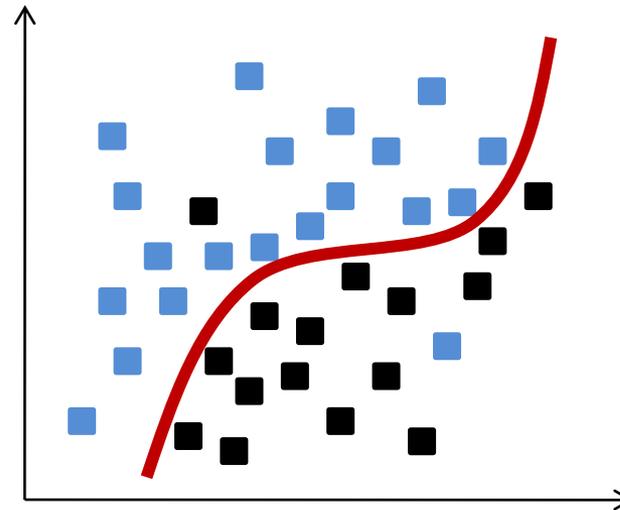
direction and strength of influence of
the first variable on the log-odds ratio
(probability of positive class)

- Moreover, **monotonicity** can easily be assured by fixing the sign of regression coefficients: If a variable increases, then the probability of the positive class must only increase (decrease)!
- *this is crucial in many applications (e.g., medicine)*
- *violation of monotonicity may often lead to the refusal of a model*

- A disadvantage of logistic regression is a **lack of flexibility**:
In many applications, the assumption of a **linear dependency** (between predictor variables and log-odds ratio), and hence a **linear decision boundary** in the instance space, is not valid!



linear decision boundary



nonlinear decision boundary

- A disadvantage of logistic regression is a **lack of flexibility**: In many applications, the assumption of a **linear dependency** (between predictor variables and log-odds ratio), and hence a **linear decision boundary** in the instance space, is not valid!
- Key question addressed in this paper:

How to increase the flexibility of logistic regression without losing its advantages of interpretability and monotonicity?

- Our general idea is to replace the linear model by the **Choquet integral** as a more flexible operator for aggregating the input attributes!

Logistic

$$\mathbf{P}(y = 1 | \mathbf{x}) = \left(1 + \exp \left(-w_0 - \mathbf{w}^\top \mathbf{x} \right) \right)^{-1}$$

Choquistic

$$\mathbf{P}(y = 1 | \mathbf{x}) = \left(1 + \exp \left(-\gamma (\mathcal{C}_\mu(\mathbf{x}) - \beta) \right) \right)^{-1}$$

Choquet integral of
(normalized) attribute values

- It can be shown that, by choosing the parameters in a proper way, logistic regression is indeed a **special case of Choquistic regression**.

Choquistic Regression: Interpretation

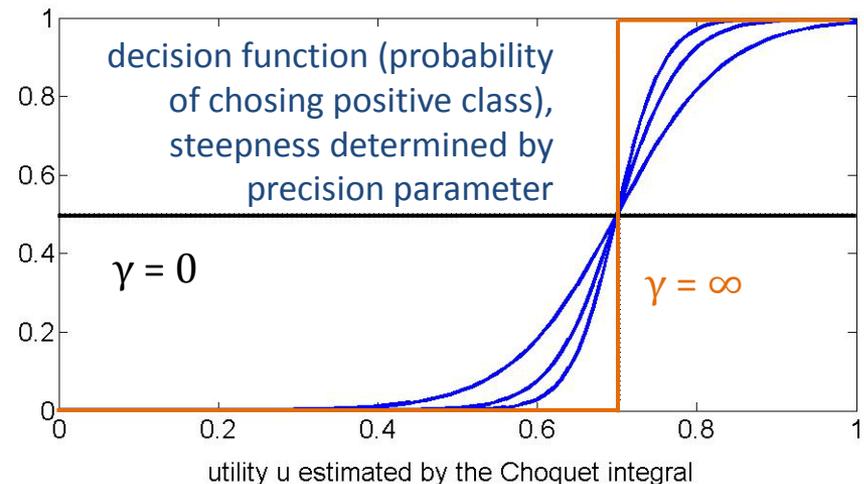
Interpretation of choquistic regression as a **two-stage process**:

- (1) a (latent) utility degree $u = \mathcal{C}_\mu(\mathbf{x}) \in [0, 1]$ is determined by the Choquet integral
- (2) a discrete choice is made by thresholding u “probabilistically” at β

Probabilistic thresholding:

$$\mathbf{P}(y = 1) = \frac{1}{1 + \exp(-\gamma(\mathcal{C}_\mu(\mathbf{x}) - \beta))}$$

↑ precision of the model ↑ utility threshold



Discrete Choquet Integral: A Brief Reminder



A **fuzzy measure** on $C = \{c_1, c_2, \dots, c_m\}$ is a set function $\mu : 2^C \rightarrow [0, 1]$ which is

- monotonic: $\mu(A) \leq \mu(B)$ for $A \subseteq B \subseteq C$
- normalized: $\mu(\emptyset) = 0$ and $\mu(C) = 1$

The **discrete Choquet integral** of $f : C \rightarrow \mathbb{R}_+$ with respect to μ is defined as follows:

$$\mathcal{C}_\mu(f) = \sum_{i=1}^m (f(c_{(i)}) - f(c_{(i-1)})) \cdot \mu(A_{(i)}) ,$$

where (\cdot) is a permutation of $\{1, \dots, m\}$ such that $0 \leq f(c_{(1)}) \leq f(c_{(2)}) \leq \dots \leq f(c_{(m)})$, and $A_{(i)} = \{c_{(i)}, \dots, c_{(m)}\}$.

In our case, $f(c_i) = x_i$ is the value of the i -th variable.

- The fuzzy measure μ specifies the **importance** of subsets of predictor variables, i.e., their influence on the probability of the positive class.
- Due to the non-additivity of this measure, it becomes possible to model **interaction effects**, thereby expressing complementarity and redundancy of variables.

*For example, what is the **joint effect** of {smoking, age} on the probability of cancer, as opposed to the sum of their individual influences?*

- Formally, measures like **Shapley index** and **interaction index** can be used, respectively, to quantify the importance of individual and the interaction between different variables.
- **Monotonicity** is obviously assured by the Choquet integral, too.

- We need to identify the following model parameters:
 - the fuzzy measure μ
 - the utility threshold β
 - the precision parameter γ
- The fuzzy measure, in its most general form, has a number of parameters which is exponential in the number of attributes
→ critical from a computational complexity point of view
- Again, we follow a **Maximum Likelihood** (ML) approach; the Choquet integral is expressed in terms of its **Möbius transform**:

$$\mathcal{C}_\mu(f) = \sum_{T \subseteq C} m(T) \times \min_{c_i \in T} f(c_i) .$$

Choquistic Regression: Parameter Estimation

- ML estimation leads to a **constrained optimization problem**:

$$\min_{\mathbf{m}, \gamma, \beta} \gamma \sum_{i=1}^n (1 - y^{(i)}) (\mathcal{C}_{\mathbf{m}}(\mathbf{x}^{(i)}) - \beta) + \sum_{i=1}^n \log \left(1 + \exp(-\gamma (\mathcal{C}_{\mathbf{m}}(\mathbf{x}^{(i)}) - \beta)) \right)$$

subject to:

$$\left. \begin{array}{l} 0 \leq \beta \leq 1 \\ 0 < \gamma \end{array} \right\} \begin{array}{l} \text{conditions on utility} \\ \text{threshold and precision} \end{array}$$

$$\left. \begin{array}{l} \sum_{T \subseteq C} \mathbf{m}(T) = 1 \\ \sum_{B \subseteq A \setminus \{c_i\}} \mathbf{m}(B \cup \{c_i\}) \geq 0 \quad \forall A \subseteq C, \forall c_i \in C \end{array} \right\} \begin{array}{l} \text{normalization and} \\ \text{monotonicity of the} \\ \text{fuzzy measure} \end{array}$$

→ solution with sequential quadratic programming

- Experimental comparison with monotone logistic regression
- Collection of data sets for which monotonicity is a plausible assumption
- Classification error determined by means of cross validation

data set	logistic	choquistic
ESL	0.0621 ± 0.0096	0.0547 ± 0.0105
ERA	0.2849 ± 0.0140	0.2756 ± 0.0170
LEV	0.1669 ± 0.0134	0.1340 ± 0.0115
DBS	0.1443 ± 0.0371	0.1560 ± 0.0405
CPU	0.0400 ± 0.0093	0.0119 ± 0.0138
CEV	0.1883 ± 0.0066	0.0346 ± 0.0076
CYD-1	0.1254 ± 0.0074	0.0729 ± 0.0066
CYD-2	0.2004 ± 0.0091	0.0717 ± 0.0078
CYD-3	0.1512 ± 0.0238	0.0762 ± 0.0163
CYD-4	0.1289 ± 0.0253	0.0496 ± 0.0201
CYD-5	0.1242 ± 0.0099	0.0204 ± 0.0057
CYD-6	0.1604 ± 0.0085	0.0383 ± 0.0083
CYD-7	0.1958 ± 0.0207	0.0646 ± 0.0089

Main results

- Choquistic regression achieves consistent gains
- Higher interaction between variables tends to come with higher gain

- We introduced a new method called **choquistic regression**, a generalization of conventional logistic regression for **binary classification**.
- **Choquistic regression**
 - combines **probabilistic modeling** underlying logistic regression with the advantages of the **Choquet integral as a flexible aggregation operator**, notably its capability to capture **interactions between predictor variables**;
 - thereby, it becomes possible to **increase flexibility** while preserving core features of logistic regression, namely **interpretability and monotonicity**.
- First **experimental results** confirm advantages of choquistic regression in terms of predictive accuracy.
- **Ongoing work:** Restriction to k-additive measures, for a properly chosen k
 - full flexibility is normally not needed and may even lead to overfitting the data
 - advantages from a computational point of view
 - key question: how to find a suitable k in an efficient way?

Back up (Influence of precision parameter)

